

New Affine Invariant Ensemble Samplers

and their dimensional scaling

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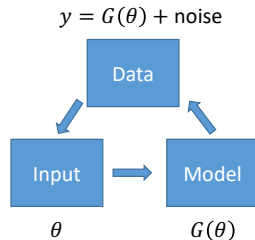
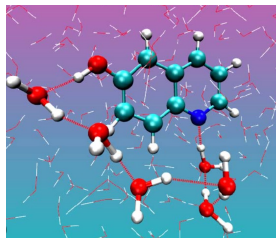
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Sampling and MCMC

Sampling from probability distributions $\pi \propto \exp(-V)$ is a classical and fundamental challenge in scientific computing and statistics

Wide applications in

- ▶ Statistical physics
- ▶ Bayes inverse problems
- ▶ Uncertainty quantification
- ▶ Filtering
- ▶ ...



Markov chain Monte Carlo (MCMC) has been the workhorse for sampling problems

at a discrete m -step: $\mathbf{x}(m+1) = R(\mathbf{x}(m), \pi)$

where R is some random mapping. As $m \rightarrow \infty$, $\mathbf{x}(m)$ approximately draws from π

Anisotropy and affine invariance [Goodman, Weare 2010]

Probability distributions can be **highly anisotropic or ill-conditioned**, such as due to hierarchical Bayes or multiple scales in physical fields

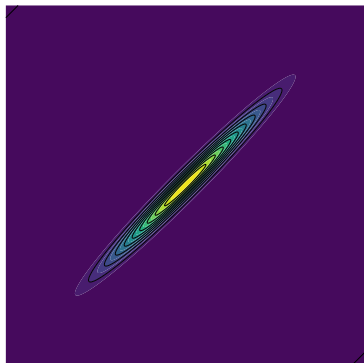


Figure: $\exp\left(-\frac{(x_1-x_2)^2}{2\epsilon} - \frac{(x_1+x_2)^2}{2}\right)$

- ▶ Ordinary MCMC may suffer from anisotropy
- ▶ **Affine invariance**: for any invertible affine transformation $\mathbf{y} = \phi(\mathbf{x}) = A\mathbf{x} + \mathbf{b}$

$$\mathbf{y}(m+1) = R(\mathbf{y}(m), \phi \# \pi)$$

Convergence behavior of $\mathbf{x}(m)$ toward π matches that of $\mathbf{y}(m)$ toward $\phi \# \pi$

- ▶ Concept motivated by affine invariance of optimization algorithms: Nelder-Mead simplex [Nelder, Mead 1965] and Newton

Ensemble affine invariant sampler [Goodman, Weare 2010]

Affine invariance in MCMC is widely achieved by ensemble samplers

Ensemble samplers: at a discrete m -step

$$(\mathbf{x}_1(m+1), \mathbf{x}_2(m+1), \dots, \mathbf{x}_N(m+1)) = R(\mathbf{x}_1(m), \mathbf{x}_2(m), \dots, \mathbf{x}_N(m), \pi)$$

Target distribution is π^N in the product space

Affine invariance: for any invertible affine transformation $\mathbf{y} = \phi(\mathbf{x}) = A\mathbf{x} + \mathbf{b}$

$$(\mathbf{x}_1, \dots, \mathbf{x}_N) \xrightarrow{\phi} (\mathbf{y}_1, \dots, \mathbf{y}_N) = (A\mathbf{x}_1 + \mathbf{b}, \dots, A\mathbf{x}_N + \mathbf{b}),$$

it holds that

$$(\mathbf{y}_1(m+1), \mathbf{y}_2(m+1), \dots, \mathbf{y}_N(m+1)) = R(\mathbf{y}_1(m), \mathbf{y}_2(m), \dots, \mathbf{y}_N(m), \phi\#\pi)$$

Ensemble affine invariant sampler [Goodman, Weare 2010]

Affine invariance in MCMC is widely achieved by ensemble samplers

Stretch move sampler

- ▶ Ensemble at step $m \in \mathbb{N}$ ($\mathbf{x}_1(m), \mathbf{x}_2(m), \dots, \mathbf{x}_N(m)$)
- ▶ **Stretch move** randomly selects two $\mathbf{x}_i(m)$ and $\mathbf{x}_j(m)$

$$\tilde{\mathbf{x}}_i(m+1) = \mathbf{x}_j(m) + Z(\mathbf{x}_i(m) - \mathbf{x}_j(m))$$

where density of Z satisfies $g(z) \propto \frac{1}{\sqrt{z}}$ for $z \in [a^{-1}, a]$; recommended $a = 2$

- ▶ Accept this proposal with probability

$$\min \left\{ 1, Z^{d-1} \frac{\pi(\tilde{\mathbf{x}}_i(m+1))}{\pi(\mathbf{x}_i(m))} \right\}$$

- ▶ Affine invariant since based on relative locations
- ▶ Used routinely in Bayesian applications: `emcee` package (>10k citations)

[Foreman-Mackey, Hogg, Lang, Goodman 2013]

Challenges in high dimensions

Affine-invariant ensemble samplers are reported to behave well for moderate dimensions but **suffer from higher dimensions** (e.g. $d \geq 50$) [Huijser, Goodman, Brewer 2015]

“Ensemble methods are doomed to fail in high dimensions” [Carpenter 2017 (blog)]

- ▶ high-dimensional distributions typically concentrate on thin shells
- ▶ interpolation or extrapolation between two points in the stretch move—as well as in many other ensemble samplers—is unlikely to fall within this shell
- ▶ small stepsizes must be used!
- ▶ the samplers effectively “devolve into random walks with poorly biased directional choices”

This work: New affine invariant ensemble samplers and their dimensional scaling

1 An ensemble side move sampler with better proposal in high dimensions

(derivative-free affine invariant samplers)

- ▶ random walk scaling of stepsize $d^{-1/2}$ for Gaussian targets
- ▶ outperform stretch move by a factor of two or more in autocorrelation time

2 Affine invariant ensemble Hamiltonian Monte Carlo

(derivative-based affine invariant samplers)

- ▶ better scaling of stepsize $d^{-1/4}$ for Gaussian targets
- ▶ can outperform derivative-free ensemble samplers and HMC by order of magnitude

Basic ensemble side move sampler (parallel version in later slides)

Side move sampler

- ▶ Ensemble at step $m \in \mathbb{N}$: $(\mathbf{x}_1(m), \mathbf{x}_2(m), \dots, \mathbf{x}_N(m))$
- ▶ **Side move** randomly selects one $\mathbf{x}_i(m)$ and two distinct $\mathbf{x}_j(m), \mathbf{x}_k(m) \neq \mathbf{x}_i(m)$

$$\tilde{\mathbf{x}}_i(m+1) = \mathbf{x}_i(m) + \sigma(\mathbf{x}_j(m) - \mathbf{x}_k(m))\xi, \quad \xi \sim \mathcal{N}(0, 1)$$

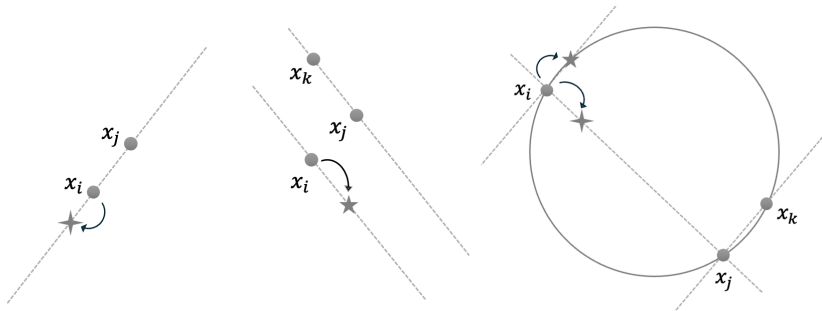
- ▶ Accept this proposal with probability

$$\min \left\{ 1, \frac{\pi(\tilde{\mathbf{x}}_i(m+1))}{\pi(\mathbf{x}_i(m))} \right\}$$

Connected to existing samplers: differential evolution and walk move (later slides)

Illustrations of stretch and side moves

- 1 Left: stretch move to a four-pointed star
- 2 Middle: side move to a five-pointed star



- 3 Right: one demo for both moves, for a ring-shaped distribution

Intuition: Side move may align better with the tangential directions

Connection to existing ensemble samplers

1 Connection to differential evolution MCMC [Ter Braak 2006]

$$\tilde{\mathbf{x}}_i(m+1) = \mathbf{x}_i(m) + \gamma(\mathbf{x}_j(m) - \mathbf{x}_k(m)) + \sigma\xi$$

where γ, σ are scalars and $\xi \sim \mathcal{N}(0, I_{d \times d})$

2 Connection to walk move [Goodman, Weare 2010]: select a subset S of particles (with mean \mathbf{m}_S) different from $\mathbf{x}_i(m)$. With i.i.d. $\xi_j \in \mathcal{N}(0, 1)$, the proposal is

$$\tilde{\mathbf{x}}_i(m+1) = \mathbf{x}_i(m) + \frac{1}{\sqrt{|S|}} \sum_{j \in S} (\mathbf{x}_j(m) - \mathbf{m}_S) \xi_j$$

When $|S| = 2$, walk move is equivalent to side move with a specific step size since

$$\frac{1}{\sqrt{|S|}} \sum_{j \in S} (\mathbf{x}_j(m) - \mathbf{m}_S) \xi_j = \frac{1}{2\sqrt{2}} (\mathbf{x}_j(m) - \mathbf{x}_k(m)) (\xi_j - \xi_k)$$

for $S = \{\mathbf{x}_j, \mathbf{x}_k\}$ and $\xi_j - \xi_k \sim \mathcal{N}(0, 2)$

Analysis of dimensional scaling for Gaussian targets

Proposition

Assume that $\mathbf{x}_i, \mathbf{x}_j, \mathbf{x}_k$ are independent draws from $\pi(\mathbf{x}) \propto \exp(-\frac{1}{2}\mathbf{x}^T \mathbf{x})$, $\mathbf{x} \in \mathbb{R}^d$:

- For side move, $\tilde{\mathbf{x}}_i(m+1) = \mathbf{x}_i(m) + \sigma(\mathbf{x}_j(m) - \mathbf{x}_k(m))\xi$, if $\sigma = \frac{\alpha}{\sqrt{d}}$, then

$$\lim_{d \rightarrow \infty} \mathbb{E}[\text{acceptance}] = \mathbb{E}[\min\{1, \exp(-\alpha^2 \xi^2 - \sqrt{2}\alpha \xi z)\}] > 0$$

where $\xi \sim \mathcal{N}(0, 1)$ is independent of $z \sim \mathcal{N}(0, 1)$

- For stretch move, $\tilde{\mathbf{x}}_i(m+1) = \mathbf{x}_j(m) + Z(\mathbf{x}_i(m) - \mathbf{x}_j(m))$, if $a = 1 + \frac{\beta}{\sqrt{d}}$, then

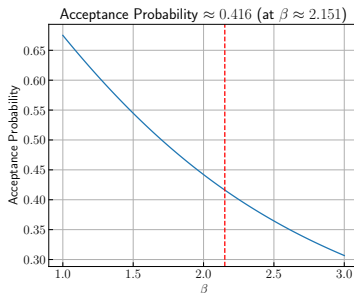
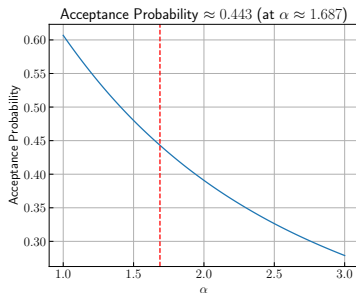
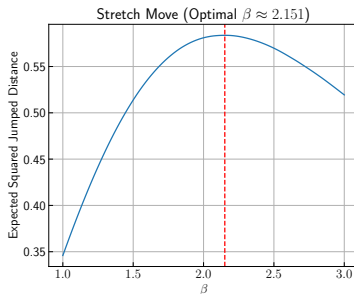
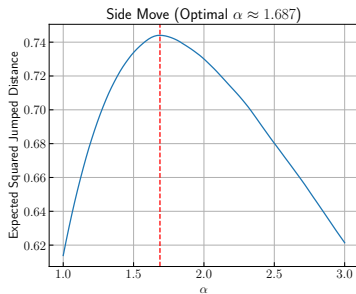
$$\lim_{d \rightarrow \infty} \mathbb{E}[\text{acceptance}] = \mathbb{E}[\min\{1, \exp\left(-\frac{3}{2}\beta^2 U^2 - \sqrt{3}\beta U z\right)\}] > 0$$

where $U \sim \text{Unif}[-1, 1]$ is independent of $z \sim \mathcal{N}(0, 1)$

- Random walk type scaling of stepsize: σ and $a - 1 \sim d^{-1/2}$

Celebrated results for single-chain MCMC [Gelman, Gilks, Roberts 1997]

Optimize the expected squared jump distance (ESJD)



ESJD in one iteration:

[Pasarica, Gelman 2010]

$$\mathbb{E}[\|\mathbf{x}_i(m+1) - \mathbf{x}_i(m)\|_2^2]$$

Optimize over stepsize:

- For side move

$$\sigma^* \approx \frac{1.687}{\sqrt{d}}$$

- For stretch move

$$a^* - 1 \approx \frac{2.151}{\sqrt{d}}$$

Optimal ESJD:

side move > stretch move

Implementation: Parallel ensemble side move

Split ensemble approach [Foreman-Mackey, Hogg, Lang, Goodman 2013]

- ▶ Divide the ensemble into two groups:

$$S^{(0)} = \{\mathbf{x}_1, \dots, \mathbf{x}_{N/2}\}, \quad S^{(1)} = \{\mathbf{x}_{N/2+1}, \dots, \mathbf{x}_N\}$$

- ▶ At each time step, for each particle in $S^{(0)}$, we randomly select two particles from the complementary set $S^{(1)}$ and perform the side move, applying the Metropolis accept-reject criterion
- ▶ Then, we follow the same procedure for particles in $S^{(1)}$, selecting particles from the complementary set $S^{(0)}$ to form the side moves and perform Metropolis
- ▶ Iterate the above two steps

This approach preserves the correct detailed balance condition

Test criterion: integrated autocorrelation time in the stationary phase

Ensemble samplers generate sequences $(\mathbf{x}_1(m), \dots, \mathbf{x}_N(m))$ for $1 \leq m \leq M$

- ▶ We estimate the observable $A = \mathbb{E}^{\mathbf{x} \sim \pi}[f(\mathbf{x})] = \int f(\mathbf{x})\pi(\mathbf{x})d\mathbf{x}$ via

$$\hat{A}_e = \frac{1}{M} \sum_{m=1}^M F(\mathbf{x}_1(m), \dots, \mathbf{x}_N(m)) = \frac{1}{M} \sum_{m=1}^M \left(\frac{1}{N} \sum_{i=1}^N f(\mathbf{x}_i(m)) \right)$$

- ▶ At the stationary phase, for large M , the variance of the estimator satisfies

$$\text{Var}(\hat{A}_e) \approx \frac{\tau_e}{M} \text{Var}^{\mathbf{x}_1, \dots, \mathbf{x}_N \sim \pi^N}[F(\mathbf{x}_1, \dots, \mathbf{x}_N)] = \frac{\tau_e}{NM} \text{Var}^{\mathbf{x} \sim \pi}[f(\mathbf{x})]$$

where τ_e is the **integrated autocorrelation time** $\tau_e = \sum_{m=-\infty}^{+\infty} \frac{C_e(m)}{C_e(0)}$ with the autocovariance function defined as

$$C_e(m) = \lim_{m' \rightarrow \infty} \text{Cov}[F(\mathbf{x}_1(m'), \dots, \mathbf{x}_N(m')), F(\mathbf{x}_1(m + m'), \dots, \mathbf{x}_N(m + m'))]$$

The **autocorrelation function** at lag m is the ratio $\frac{C_e(m)}{C_e(0)}$

Numerical experiments on high dimensional Gaussians

Target: Gaussian distributions in d dims

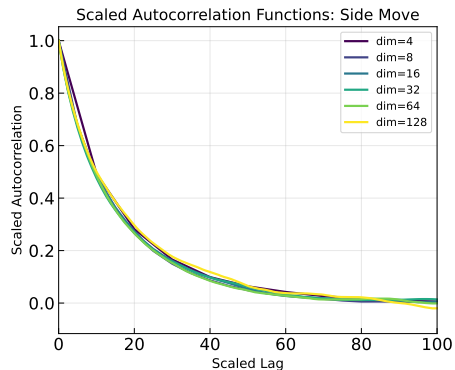
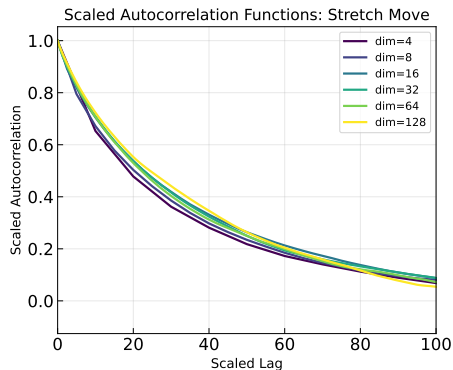


Figure: Scaled autocorrelation function for 1d position observable; Scaled lag = original lag / $\dim \times 4$

This implies the $O(d)$ scaling of autocorrelation function and autocorrelation time

Discussions

Both ensemble stretch and side moves are **affine invariant and derivative free**

- ▶ Optimal scaling of stepsize parameters $\sim d^{-1/2}$
- ▶ Side move outperforms stretch move in ESJD, and autocorrelation time in practice (more numerical examples at the end)
- ▶ Still, bottleneck of $O(d)$ autocorrelation time: intrinsic to random walk scaling

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Derivative-based samplers have been shown to scale better with dimension

- ▶ Metropolized Langevin stepsize $\sim d^{-1/3}$ [Roberts, Rosenthal 1998]
- ▶ Hamiltonian Monte Carlo stepsize $\sim d^{-1/4}$ [Beskos, Pillai, Roberts, Sanz-Serna, Stuart 2013]

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Many existing work to adapt these derivative-based samplers to anisotropy [Girolami, Calderhead 2011], [Martin, Wilcox, Burstedde, Ghattas 2012], [Greengard 2015], [Simsekli, Badeau, Cemgil, Richard 2016], [Leimkuhler, Matthews, Weare 2018], [Kleppe 2019], [Garbuno-Inigo, Hoffmann, Li, Stuart 2020], [Garbuno-Inigo, Nüsken, Reich 2020], [Hoffman, Sountsov 2022], [Tran, Kleppe 2024], etc.

This work: New affine invariant ensemble samplers and their dimensional scaling

1 An ensemble side move sampler with better proposal in high dimensions (derivative-free affine invariant samplers)

- ▶ random walk scaling of stepsize $d^{-1/2}$ for Gaussian targets
- ▶ outperform stretch move by a factor of two or more in autocorrelation time

2 Affine invariant ensemble Hamiltonian Monte Carlo (derivative-based affine invariant samplers)

- ▶ better scaling $d^{-1/4}$ of stepsize for Gaussian targets
- ▶ can outperform derivative-free ensemble samplers and HMC by order of magnitude

Hamiltonian Monte Carlo (HMC)

For $\pi \propto \exp(-V)$, HMC augments it to $\tilde{\pi} \propto \exp(-V(\mathbf{x}) - \frac{1}{2}\mathbf{p}^T M^{-1}\mathbf{p})$

- ▶ Hamiltonian dynamics $\frac{d\mathbf{x}}{dt} = M^{-1}\mathbf{p}$, $\frac{d\mathbf{p}}{dt} = -\nabla V(\mathbf{x})$ keep $\tilde{\pi}$ invariant
- ▶ Discretization: L_h is one leapfrog with step size h , i.e., $(\mathbf{x}_h, \mathbf{p}_h) = L_h(\mathbf{x}, \mathbf{p})$ with

$$\mathbf{p}_{h/2} = \mathbf{p} - \frac{h}{2}\nabla V(\mathbf{x}), \quad \mathbf{x}_h = \mathbf{x} + hM^{-1}\mathbf{p}_{h/2}, \quad \mathbf{p}_h = \mathbf{p}_{h/2} - \frac{h}{2}\nabla V(\mathbf{x}_h)$$

- ▶ P is the momentum flip operator such that $P(\mathbf{x}, \mathbf{p}) = (\mathbf{x}, -\mathbf{p})$

HMC algorithm: alternating between two steps

- 1 Sample a momentum $\mathbf{p} \sim \mathcal{N}(0, M)$
- 2 Propose an update $(\tilde{\mathbf{x}}, \tilde{\mathbf{p}}) = PL_h^n(\mathbf{x}, \mathbf{p})$ and accept with probability

$$\text{prob} = \min \left\{ 1, \exp \left(-V(\tilde{\mathbf{x}}) - \frac{1}{2}\tilde{\mathbf{p}}^T M^{-1}\tilde{\mathbf{p}} + V(\mathbf{x}) + \frac{1}{2}\mathbf{p}^T M^{-1}\mathbf{p} \right) \right\}$$

Importance of tuning of M, h, n : NUTS [Hoffman, Gelman 2014], Stan [Carpenter et al. 2017], etc.

First attempt: Covariance preconditioning using mass matrices

Attempt: The simplest idea is to set M^{-1} to be a covariance matrix

For the sake of parallel algorithm, we use split ensembles with two groups

$$S^{(0)} = \{\mathbf{x}_1, \dots, \mathbf{x}_{N/2}\}, \quad S^{(1)} = \{\mathbf{x}_{N/2+1}, \dots, \mathbf{x}_N\}$$

The joint distribution ($\text{Cov}_{S^{(0)}}, \text{Cov}_{S^{(1)}}$): empirical covariance matrices of $S^{(0)}, S^{(1)}$

$$\exp\left(-\sum_{i=1}^{N/2} (V(\mathbf{x}_i) + \frac{1}{2} \mathbf{p}_i^T \text{Cov}_{S^{(1)}} \mathbf{p}_i) - \sum_{i=N/2+1}^N (V(\mathbf{x}_i) + \frac{1}{2} \mathbf{p}_i^T \text{Cov}_{S^{(0)}} \mathbf{p}_i)\right)$$

does not preserve the correct marginal on \mathbf{x}_i

Resolution: need to add $-\frac{N}{4} \log \det \text{Cov}_{S^{(0)}} - \frac{N}{4} \log \det \text{Cov}_{S^{(1)}}$ to the potential

Issue: particles in each group must be accepted or rejected at the same time

Better idea: Antisymmetric preconditioning

Consider the simple joint distribution

$$\exp \left(- \sum_{i=1}^{N/2} \left(V(\mathbf{x}_i) + \frac{1}{2} \mathbf{p}_i^T \mathbf{p}_i \right) - \sum_{i=N/2+1}^N \left(V(\mathbf{x}_i) + \frac{1}{2} \mathbf{p}_i^T \mathbf{p}_i \right) \right)$$

Key observation: The following preconditioned Hamiltonian dynamics

$$\frac{d\mathbf{x}_i}{dt} = \mathbf{B} \mathbf{p}_i, \quad \frac{d\mathbf{p}_i}{dt} = -\mathbf{B}^T \nabla V(\mathbf{x}_i)$$

leaves the distribution invariant

This is an antisymmetric preconditioning of gradient flow

$$\begin{bmatrix} \frac{d\mathbf{x}_i}{dt} \\ \frac{d\mathbf{p}_i}{dt} \end{bmatrix} = \underbrace{\begin{bmatrix} 0 & \mathbf{B} \\ -\mathbf{B}^T & 0 \end{bmatrix}}_{\text{antisymmetric}} \begin{bmatrix} \nabla V(\mathbf{x}_i) \\ \mathbf{p}_i \end{bmatrix}$$

Used in [Leimkuhler, Matthews, Weare 2018] for preconditioning underdamped Langevin

Affine invariant ensemble HMC

- **Fix particles in $S^{(1)}$, and update particles in $S^{(0)}$:** sample $\mathbf{p}_i \sim \mathcal{N}(0, I_{N/2 \times N/2})$, run n steps of leapfrog of the dynamics and apply Metropolis

$$\frac{d\mathbf{x}_i}{dt} = \mathbf{B}_{S^{(1)}} \mathbf{p}_i, \quad \frac{d\mathbf{p}_i}{dt} = -\mathbf{B}_{S^{(1)}}^T \nabla V(\mathbf{x}_i), \quad 1 \leq i \leq N/2$$

with $(\mathbf{m}_{S^{(1)}})$ is the mean of all particles in $S^{(1)}$

$$\mathbf{B}_{S^{(1)}} = \frac{1}{\sqrt{N/2}} [\mathbf{x}_{N/2+1} - \mathbf{m}_{S^{(1)}}, \dots, \mathbf{x}_N - \mathbf{m}_{S^{(1)}}] \in \mathbb{R}^{d \times N/2}$$

- **Fix particles in $S^{(0)}$, and update particles in $S^{(1)}$ similarly:** dynamic is now

$$\frac{d\mathbf{x}_i}{dt} = \mathbf{B}_{S^{(0)}} \mathbf{p}_i, \quad \frac{d\mathbf{p}_i}{dt} = -\mathbf{B}_{S^{(0)}}^T \nabla V(\mathbf{x}_i), \quad N/2 + 1 \leq i \leq N$$

with $(\mathbf{m}_{S^{(0)}})$ is the mean of all particles in $S^{(0)}$

$$\mathbf{B}_{S^{(0)}} = \frac{1}{\sqrt{N/2}} [\mathbf{x}_1 - \mathbf{m}_{S^{(0)}}, \dots, \mathbf{x}_{N/2} - \mathbf{m}_{S^{(0)}}] \in \mathbb{R}^{d \times N/2}$$

- Iterate the above two steps

Dimensional scaling analysis of the affine invariant ensemble HMC

Proposition: Assume that \mathbf{x}_i are independent draws from $\pi(\mathbf{x}) \propto \exp(-\frac{1}{2}\mathbf{x}^T \mathbf{x})$, $\mathbf{x} \in \mathbb{R}^d$. If we take $h = \alpha d^{-1/4}$, $n = T/h$ and $\lim_{d \rightarrow \infty} \frac{d}{N/2} = \rho \in [0, 1)$, then

$$\lim_{d \rightarrow \infty} \mathbb{E}[\text{acceptance}] = \mathbb{E}[\min\{1, \exp(\mathcal{N}(\alpha^4 \mu_\rho, \alpha^4 \sigma_\rho))\}] > 0$$

where $\mathcal{N}(\alpha^4 \mu_\rho, \alpha^4 \sigma_\rho)$ is Gaussian with

$$\mu_\rho = -\frac{1}{32} \int \lambda^4 \sin^2(\sqrt{\lambda} T) d\nu_\rho(\lambda), \quad \sigma_\rho = \frac{1}{16} \int \lambda^6 \sin^2(\sqrt{\lambda} T) d\nu_\rho(\lambda)$$

For $\rho \in [0, 1)$, $d\nu_\rho(\lambda) = \frac{1}{2\pi\rho\lambda} \sqrt{(c-\lambda)(\lambda-b)} \chi_{[b,c]}(\lambda) d\lambda$, where $b = (1 - \sqrt{\rho})^2$, $c = (1 + \sqrt{\rho})^2$ and $\chi_{[b,c]}$ is the characteristic function of $[b, c]$

Moreover, the **expected squared jumped distance (ESJD)** for one iteration satisfies

$\lim_{d \rightarrow \infty} \frac{1}{d} \text{ESJD} = c \in (0, \infty)$, in contrast to $\lim_{d \rightarrow \infty} \text{ESJD} = c' \in (0, \infty)$ for stretch/side moves

Variant: directional side move in ensemble HMC

- **Fix particles in $S^{(1)}$, and update particles in $S^{(0)}$:** for $1 \leq i \leq N/2$, sample $p_i \sim \mathcal{N}(0, 1)$ and two particles \mathbf{x}_j and \mathbf{x}_k from $S^{(1)}$, run n steps of leapfrog of the dynamics and apply Metropolis

$$\frac{d\mathbf{x}_i}{dt} = \frac{1}{\sqrt{2d}}(\mathbf{x}_j - \mathbf{x}_k)p_i, \quad \frac{dp_i}{dt} = -\frac{1}{\sqrt{2d}}(\mathbf{x}_j - \mathbf{x}_k)^T \nabla V(\mathbf{x}_i)$$

- **Fix particles in $S^{(0)}$, and update particles in $S^{(1)}$ similarly**
- Iterate the above two steps

We call the algorithm **Hamiltonian side move**; the previous one **Hamiltonian walk move**

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We call the algorithm **Hamiltonian side move**; the previous one **Hamiltonian walk move**

Proposition: Assume that \mathbf{x}_i are independent draws from $\pi(\mathbf{x}) \propto \exp(-\frac{1}{2}\mathbf{x}^T \mathbf{x})$, $\mathbf{x} \in \mathbb{R}^d$. Let $h = \alpha < 2$ independent from d , then for one iteration:

$$\lim_{d \rightarrow \infty} \mathbb{E}[\text{acceptance}] > 0, \quad \lim_{d \rightarrow \infty} \text{ESJD} = c' \in (0, \infty)$$

Hamiltonian side move has the same scaling as the derivative-free side move

Numerical experiments: anisotropic Gaussian

Target: Gaussian distributions in d dims with condition number $\kappa = 10^3$

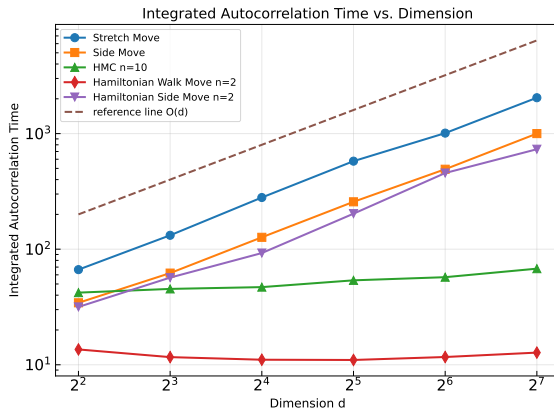


Figure: Autocorrelation time (1D position observable) versus dimension. For HMC and affine invariant HMC (Hamiltonian walk and side moves), the total integration time is $T = 1$ and n is the number of leapfrog steps

Numerical experiments: anisotropic Gaussian

Target: Gaussian distributions in $d = 128$ dims with condition number $\kappa = 10^3$

	acceptance rate	autocorrelation time τ_e or τ_s	func eval per iter	grad eval per iter
Stretch move	0.45	2043.6	1	0
Side move	0.45	1000.1	1	0
HMC: $n = 10$	0.57	67.8	1	11
HMC: $n = 2$	0.00	—	1	3
Hamiltonian walk move: $n = 10$	0.98	10.5	1	11
Hamiltonian walk move: $n = 2$	0.61	12.7	1	3
Hamiltonian side move: $n = 10$	1.00	898.2	1	11
Hamiltonian side move: $n = 2$	0.98	732.3	1	3

Table: For HMC and affine invariant HMC (Hamiltonian walk and side moves), the total integration time is $T = 1$ and n is the number of leapfrog steps

Numerical experiments: physical fields (discretized by finite difference with d grid points)

$$\pi(u) \propto \exp \left(- \int_0^1 \frac{1}{2} (\partial_x u(x))^2 + V(u(x)) dx \right), \quad V(u) = (1 - u^2)^2$$

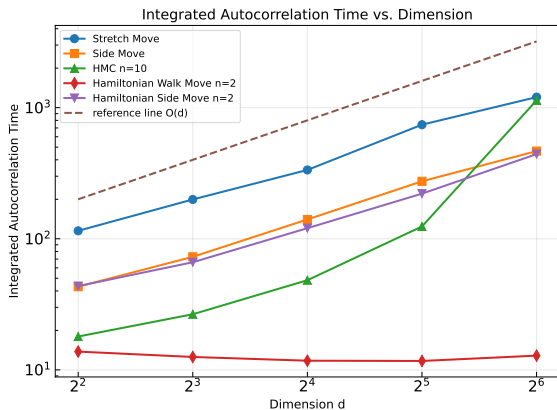


Figure: Autocorrelation time (integral observable $\int_0^1 u(x) dx$) versus dimension. For HMC and affine invariant HMC (Hamiltonian walk and side moves), the total integration time is $T = 1$ and n is the number of leapfrog steps

Discussions

- ▶ **Side move**: more favorable high dimensional ensemble proposal than stretch move
Nevertheless random walk scaling bottleneck applies
- ▶ **Affine invariant HMC**: derivative-based ensemble sampler
Antisymmetric preconditioning and much better scaling for stepsize $h \sim d^{-1/4}$
Further adaptation of stepsize and integration length?
- ▶ **Affine invariant Hamiltonian side move**: random walk scaling but slightly better than side move per iteration (further potential adaptation of derivative-free side moves?)
- ▶ Extension of analysis beyond Gaussians and stationary phases
- ▶ Broader invariance? Diffemorphism invariance of Fisher-Rao gradient flows and variational inference [Chentsov 1985], [Chen, Huang, Huang, Reich, Stuart 2023, 2024], etc.

Thank you!

Back-Up Slides

Numerical experiments: physical fields (discretized by finite difference with d grid points)

$$\pi(u) \propto \exp \left(- \int_0^1 \frac{1}{2} (\partial_x u(x))^2 + V(u(x)) dx \right), \quad V(u) = (1 - u^2)^2$$

	acceptance rate	autocorrelation time τ_e or τ_s	func eval per iter	grad eval per iter
Stretch move	0.44	3021.3	1	0
Side move	0.44	1398.3	1	0
HMC: $n = 10$	0.00	—	1	11
HMC: $n = 2$	0.00	—	1	3
Hamiltonian walk move: $n = 10$	0.98	11.2	1	11
Hamiltonian walk move: $n = 2$	0.59	14.9	1	3
Hamiltonian side move: $n = 10$	1.00	902.8	1	11
Hamiltonian side move: $n = 2$	0.98	770.8	1	3

Table: For HMC and affine invariant HMC (Hamiltonian walk and side moves), the total integration time is $T = 1$ and n is the number of leapfrog steps

Numerical experiments: ring-shaped distributions in d dimensions

$$\pi(\mathbf{x}) \propto \exp\left(-(\|\mathbf{x}\|_2^2 - 1)^2/l^2\right), l = 0.25, d = 50$$

	acceptance rate	autocorrelation time τ_e or τ_s	func eval per iter	grad eval per iter
Stretch move	0.29	2435.4	1	0
Side move	0.45	355.4	1	0
HMC: $n = 10$	0.69	20.8	1	11
HMC: $n = 2$	0.00	—	1	3
Hamiltonian walk move: $n = 10$	0.99	10.7	1	11
Hamiltonian walk move: $n = 2$	0.72	11.9	1	3
Hamiltonian side move: $n = 10$	1.00	354.8	1	11
Hamiltonian side move: $n = 2$	0.98	309.7	1	3

Table: For HMC and affine invariant HMC (Hamiltonian walk and side moves), the total integration time is $T = 1$ and n is the number of leapfrog steps